

Probabilistic Method and Random Graphs

Lecture 10. Second Moment Method and Lovász Local Lemma

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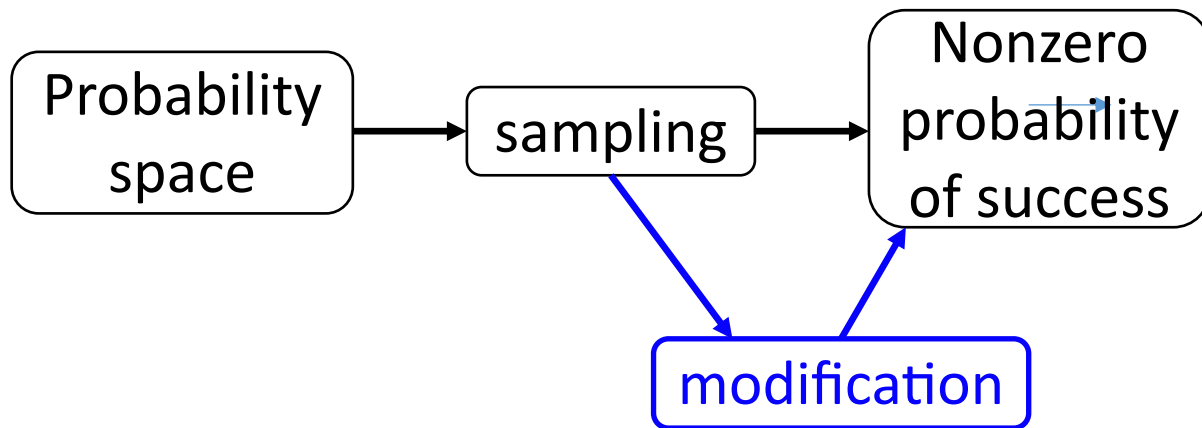
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¹The slides are mainly based on Chapter 6 of Probability and Computing.

Comments, questions, or suggestions?

Recap of Lecture 9

- Derive a deterministic algorithm from expectation argument
- Markov's Ine.: graphs with arbitrarily big girth and chro. number



First
Moment
method

Recap of Lecture 9

- Chebyshev's Ine.: $\Pr(|X - \mathbb{E}[X]| \geq a) \leq \frac{\text{Var}[X]}{a^2}$
 - $\Pr(X = 0) \leq \Pr(|X - \mathbb{E}[X]| \geq \mathbb{E}[X]) \leq \frac{\text{Var}[X]}{\mathbb{E}[X^2]} \leq \frac{\text{Var}[X]}{(\mathbb{E}[X])^2}$
 - When $X \geq 0$, $\Pr(X > 0) > \frac{(\mathbb{E}[X])^2}{\mathbb{E}[X^2]}$
 - $\Pr(X > \theta \mathbb{E}[X]) \geq \frac{(1-\theta)^2 (\mathbb{E}[X])^2}{\text{Var}[X] + (1-\theta)^2 (\mathbb{E}[X])^2}$
 $\geq (1 - \theta)^2 \frac{(\mathbb{E}[X])^2}{\mathbb{E}[X^2]}, \theta \in (0,1)$
 - Application: Distinct Subset sum problem
- } Second
moment
method

Main Probabilistic Methods

- Counting argument
- First-moment method
- **Second-moment method**
- Lovasz local lemma

Application: threshold function

- Consider a property P of random graph $G_{n,p}$

- Threshold function $t(n)$ for P is such that

$$\lim_{n \rightarrow \infty} \Pr(G_{n,p} \text{ has } P) = \begin{cases} 0 & \text{if } p = o(t(n)) \\ 1 & \text{if } p = \omega(t(n)) \end{cases}$$

- **Example** (clique number $c(G)$: max clique size)

- $P: c(G) \geq 4$

- $t(n) = n^{-\frac{2}{3}}$ is its threshold function

Proof: when $p = o(n^{-\frac{2}{3}})$

- S : a 4-subset of the n vertices
- X_S : indicator of whether S spans a clique
- $X = \sum_S X_S$: the number of 4-cliques

- $\mathbb{E}[X] = \binom{n}{4} p^6 < \frac{n^4 p^6}{24}$

- By Markov's inequality

$$\begin{aligned} \Pr(c(G) \geq 4) &= \Pr(X > 0) \\ &\leq \mathbb{E}[X] < \frac{n^4 p^6}{24} = o(1) \end{aligned}$$

Proof: when $p = \omega(n^{-\frac{2}{3}})$

- To derive $\Pr(X > 0) \rightarrow 1$
 - By Chebychev's Ineq.: $\Pr(X = 0) \leq \frac{\text{Var}[X]}{(\mathbb{E}[X])^2}$
 - Try to show $\text{Var}[X] = o(\mathbb{E}[X])^2$
- Recall $\text{Var}[X] = \sum \text{Var}[X_S] + \sum_{S \neq T} \text{Cov}(X_S, X_T)$
- X_S is an indicator $\Rightarrow \text{Var}[X_S] \leq \mathbb{E}[X_S]$
- $\text{Cov}(X_S, X_T) \leq \mathbb{E}[X_S X_T] = \Pr(X_S = 1, X_T = 1)$
 $= \mathbb{E}[X_S] \Pr(X_T = 1 | X_S = 1)$

And $\text{Cov}(X_S, X_T) = 0$ if independent

Proof: estimating the variance

- $\text{Var}[X] \leq \mathbb{E}[X] + \sum \mathbb{E}[X_S] \sum_{T \sim S} \Pr(X_T = 1 | X_S = 1)$
 $= \sum \mathbb{E}[X_S] \Delta_S$
- $\Delta_S = 1 + \sum_{|T \cap S|=2} \Pr(X_T = 1 | X_S = 1)$
 $+ \sum_{|T \cap S|=3} \Pr(X_T = 1 | X_S = 1)$
 $= 1 + \binom{n-4}{2} \binom{4}{2} p^5 + \binom{n-4}{1} \binom{4}{3} p^3$
 $= o(n^4 p^6) = o(\mathbb{E}[X])$
- $\text{Var}[X] = o(\mathbb{E}[X]^2) \Rightarrow \Pr(X = 0) \leq \frac{\text{Var}[X]}{\mathbb{E}[X]^2} = o(1)$
 $\Rightarrow \Pr(X > 0) \rightarrow 1$

Main Probabilistic Methods

- Counting argument
- First-moment method
- Second-moment method
- Lovász local lemma

Lovász local lemma: motivation

- Can we avoid all bad events?
- Given bad events A_1, A_2, \dots, A_n , is $\Pr(\cap_i \overline{A_i}) > 0$?
 - Applicable to [SAT](#), coloring, Ramsey theory...
- Two special cases
 - $\sum_i \Pr(A_i) < 1 \Rightarrow \Pr(\cap_i \overline{A_i}) \geq 1 - \sum_i \Pr(A_i) > 0$
 - Independent $\Rightarrow \Pr(\cap_i \overline{A_i}) = \prod(1 - \Pr(A_i)) > 0$
- What if *almost* independent?

Lovász local lemma: symmetric version

- Dependency graph
 - Undirected simple graph on $S = \{A_1, A_2, \dots, A_n\}$
 - A_i is independent of its non-neighborhood $S \setminus \Gamma^+(A_i)$
 - $\Gamma(A_i) \triangleq \Gamma^+(A_i) \setminus \{A_i\}$
- **Theorem:** $\Pr(\cap_i \overline{A_i}) > 0$ if
 1. $\forall i, \Pr(A_i) \leq p, |\Gamma(A_i)| \leq d$ and
 2. $4pd \leq 1$
- By Erdős&Lovász in 1973 to Erdős 60th birthday



Lovász



Erdos

Lovász local lemma: proof

- Standard trick
 - Chain rule: $\Pr(\cap_i \bar{A}_i) = \prod_{i=1}^n \Pr(\bar{A}_i | \cap_{j=1}^{i-1} \bar{A}_j)$
 - Valid only if each $\cap_{j=1}^{i-1} \bar{A}_j$ has nonzero probability
 - Hold if each term $\Pr(\bar{A}_i | \cap_{j=1}^{i-1} \bar{A}_j) > 0$
- **Claim**: for any $t \geq 0$ and $A, B_1, B_2, \dots, B_t \in S$,
 1. $\Pr(\cap_{j=1}^t \bar{B}_j) > 0$
 2. $\Pr(A | \cap_{j=1}^t \bar{B}_j) < \frac{1}{2d}$

Inductive proof of the claim

- **Basis:** $t = 0$. Both 1 and 2 of the claim hold
- **Hypothesis:** the claim holds for all $t' < t$
- **Induction**
 - For **1**, $\Pr(\bigcap_{j=1}^t \bar{B}_j)$
$$= \Pr(\bar{B}_t | \bigcap_{j=1}^{t-1} \bar{B}_j) \Pr(\bigcap_{j=1}^{t-1} \bar{B}_j) > 0$$
 - For **2**, let $\{C_1, \dots, C_x\} = \{B_1, \dots, B_t\} \cap \Gamma(A)$, and
$$\{D_1, \dots, D_y\} = \{B_1, \dots, B_t\} \setminus \Gamma(A)$$
 - $x \leq d, x + y = t$

Proof: induction for 2

- If $x = 0$, A is independent of $\{B_1, \dots, B_t\}$ and $\Pr(A | \bigcap_{j=1}^t \bar{B}_j) = \Pr(A) < \frac{1}{2d}$
- Assume $x > 0$. Then $y < t$.

$$\begin{aligned} \bullet \Pr(A | \bigcap_{j=1}^t \bar{B}_j) &= \frac{\Pr(A \cap (\bigcap_{j=1}^t \bar{B}_j))}{\Pr(\bigcap_{j=1}^t \bar{B}_j)} \\ &\leq \frac{\Pr(A \cap (\bigcap \bar{D}_j))}{\Pr((\bigcap \bar{C}_j) \cap (\bigcap \bar{D}_j))} = \frac{\Pr(A | \bigcap \bar{D}_j)}{\Pr((\bigcap \bar{C}_j) | \bigcap \bar{D}_j)} \\ &= \frac{\Pr(A)}{1 - \Pr((\bigcup C_j) | \bigcap \bar{D}_j)} < \frac{p}{1 - \frac{d}{2d}} \leq \frac{1}{2d} \end{aligned}$$

General case

Application to (k,s) -SAT

- (k,s) -CNF
 - Any clause has k literals
 - Any literal appears in at most s clauses
- Theorem: Any (k, s) -CNF is satisfiable if $s \leq \frac{1}{4} \frac{2^k}{k}$
 - Randomly assign values to the Boolean variables
 - A_i : the event that the i th clause is not satisfied
 - $\Pr(\bigcap \overline{A_i}) > 0 \Leftrightarrow$ satisfiable
 - $p = \Pr(A_i) = 2^{-k}$, $d \leq ks$
 - $s \leq \frac{1}{4} \frac{2^k}{k} \Rightarrow 4pd \leq 1 \Rightarrow \Pr(\bigcap \overline{A_i}) > 0 \Rightarrow$ satisfiable

Application to Ramsey Number $R(k)$

- Counting argument: $R(k) \geq k2^{\frac{k}{2}} \left[\frac{1}{e\sqrt{2}} + o(1) \right]$ [1947]
- Best result: $R(k) \geq k2^{\frac{k}{2}} \left[\frac{\sqrt{2}}{e} + o(1) \right]$ [1975, Spencer]
 - Randomly color edges of K_n in red/blue
 - S : a k -subset of the vertices
 - A_S : S is monochromatic
 - $p = \Pr(A_S) = 2^{1-\binom{k}{2}}$, $d \leq \binom{k}{2} \binom{n}{k-2}$
 - By Stirling's formula, $4pd \leq 1$ if $n \leq k2^{\frac{k}{2}} \left[\frac{\sqrt{2}}{e} + o(1) \right]$
- Can we say something about $R(k, t)$?

Non-symmetric LLL

- **Theorem:** $\Pr(\cap_i \overline{A_i}) > 0$ if $\forall i, \sum_{j \in \Gamma(A_i)} \Pr(A_j) < \frac{1}{4}$
 - [Spencer, 1975]
 - The sense of “local”
- Follow the proof of symmetric LLL, with induction on m to show that $\Pr(A \mid \cap_{j=1}^m \overline{B_j}) < 2\Pr(A)$
- Application to $R(k, t)$:

$$R(k, t) > t^{\frac{\binom{k}{2}-2}{k-2} + o(1)} \text{ with } k \text{ fixed and } t \rightarrow \infty$$

$$\text{Proof: } R(k, t) > t \frac{\binom{k}{2}^{-2}}{k-2} + o(1)$$

- Randomly color edges of K_n , p in red, $(1 - p)$ in blue
- S : a k -set of the vertices; T : a t -set of the vertices
- A_S : S is a red clique; B_T : T is a blue clique
- $\Pr(A_S) = p^{\binom{k}{2}}, \Pr(B_T) = (1 - p)^{\binom{t}{2}}$
- Any event has at most $\binom{t}{2} \binom{n}{k-2}$ neighbors being A_S ,
at most $\binom{n}{t}$ neighbors being B_T
- Let $p = n^{-\epsilon - \beta + \delta}$, $t = n^{\beta + \epsilon}$, $\beta = \frac{k-2}{\binom{k}{2}^{-2}}$, $0 < \delta < \epsilon$,
we have $\binom{t}{2} \binom{n}{k-2} p^{\binom{k}{2}} + \binom{n}{t} (1 - p)^{\binom{t}{2}} < \frac{1}{4}$

A stronger non-symmetric LLL

- $\Pr(\cap \overline{A}_i) > 0$ if there are $x_1, x_2, \dots, x_n \in (0,1)$ s.t.

$$\forall i, \Pr(A_i) \leq x_i \prod_{j \in \Gamma(A_i)} (1 - x_j)$$

- Similar proof, but

- Prove $\Pr(A_i | \cap_{j=1}^t \overline{B}_j) \leq x_i$

- Use chain rule to lower-bound the numerator $\Pr(\cap \overline{C}_j | \cap \overline{D}_j)$

$$\text{by } \prod_{j \in \Gamma(A_i)} (1 - x_j)$$

- Spencer, 1977

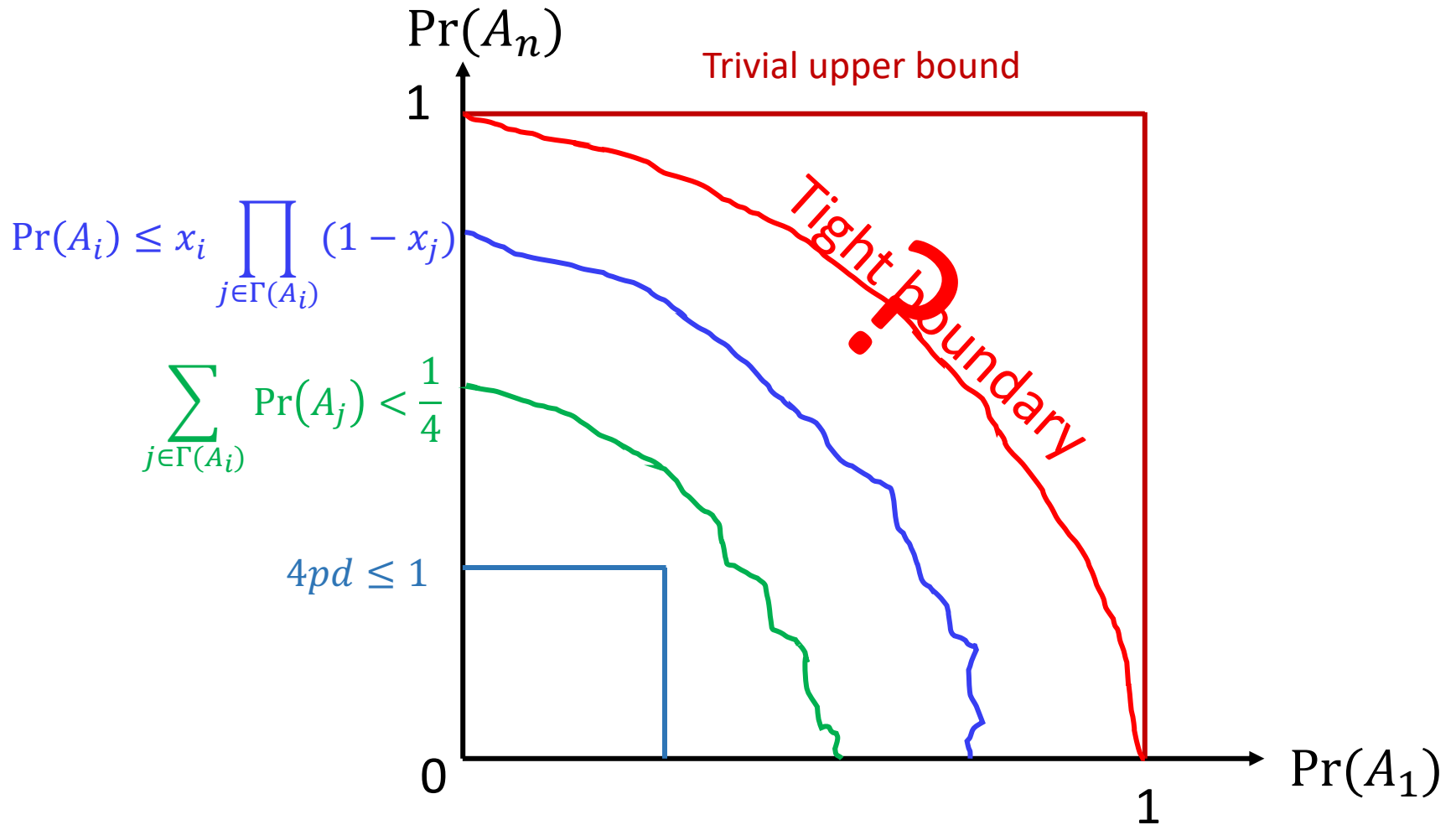
$$R(k, t) \geq c \left(\frac{t}{\ln t} \right)^{\frac{k+1}{2}} (1 - o(1))$$

- Follow the proof of $R(k, t) > t^{\frac{\binom{k}{2}-2}{k-2}+o(1)}$
 - Define events A_S and B_T for any k -set S and t -set T
 - Let $p = c_1 n^{-\beta}$, $t = c_2 n^\beta \ln n$, $x_S = (1 + \epsilon) \Pr(A_S)$
 $x_T = e^{c_3 n^\beta \ln^2 n} \Pr(A_S)$, with $\beta = \frac{2}{k+1}$, $\epsilon > 0$
 - Apply LLL
- Best until 2010
 - Bohman&Keevash: $R(k, t) \geq c \left(\frac{t}{\ln t} \right)^{\frac{k+1}{2}} (\ln t)^{\frac{1}{k-2}}$

Major open problem

- Determine $\alpha(k)$ s.t. $R(k, t) = t^{\alpha(k)+o(1)}$
- [Spencer 1975](#): $\alpha(k) \geq \frac{\binom{k}{2}-2}{k-2}$
- Spencer 1977: $\alpha(k) \geq \frac{k+1}{2} = \frac{\binom{k}{2}-1}{k-2}$
 - Best for 40+ years
 - How tight is it?
- $\alpha(k) \leq k - 1$ since $R(k, t) \leq \binom{k+t-2}{k-1}$
- Conjecture: $\alpha(k) = k - 1$
 - Yes for $k = 3$
 - Unknown for larger k

- This local lemma is so strong. Is it ultimate?



Local lemma is to determine a curve surrounding a *safe zone*.
 Safe: $\Pr(\cap_i \bar{A}_i) > 0$ any set of events with the probabilities

Tight Bound of Lovász **local** lemma

- General (Non-symmetrical) case
- $\Pr(\cap \overline{A}_i) > 0$ if

$$\forall S \in \text{ind}(G), \sum_{T \supseteq S} (-1)^{|T \setminus S|} \prod_{i \in T} p_i > 0$$

- By James B. Shearer @IBM in 1985

James Shearer



Tight Bound of Lovász **local** lemma

- Symmetrical case

- $\Pr(\bigcap \overline{A_i}) > 0$ if

$$p < \begin{cases} \frac{(d-1)^{d-1}}{d^d} & \text{when } d > 1 \\ \frac{1}{2} & \text{when } d = 1 \end{cases}$$

- **Corollary:** $\Pr(\bigcap \overline{A_i}) > 0$ if $edp \leq 1$

Application

- Any (k,s) -CNF is satisfiable if $s \leq \frac{1}{e} \frac{2^k}{k}$
 - Known: satisfiable if $s \leq \frac{1}{4} \frac{2^k}{k}$
 - Tight bound of s : $\left(\frac{2}{e} + o\left(\frac{1}{\sqrt{k}}\right)\right) \frac{2^k}{k}$
[Gebauer et al. 2011]
 - Can we efficiently find a satisfying assignment?

Algorithmic aspects

- **Like** other probabilistic methods, LLL proves existence non-constructively
- **Unlike** other probabilistic methods, LLL doesn't lead to efficient algorithms
 - Directly sampling has an *exponentially small* lower bound of success probability
 - Say, $\Pr(\cap \overline{A_i}) \geq \prod(1 - x_i)$ for general version
- Is there an efficient, constructive proof?

Constructive Lovász Local Lemma

- Initiated by Jozsef Beck in 1991
 - Under strong conditions on neighborhood size
 - In terms of coloring, SAT ...
- Breakthrough by Robin Moser & Gabor Tardos in 2009, Kashyap Kolipaka and Mario Szegedy in 2011
 - Events are generated by independent random variables
 - If Shearer's condition is met, an assignment s.t. none events occurs can be found in linear time

Josze Beck



Gabor Tardos



Mario Szegedy

The assignment algorithm

For $X \in \mathcal{X}$ do

$v_X \leftarrow$ a random evaluation of X

EndFor

While (some A occurs) do

 Arbitrarily pick an event A that occurs

 For $X \in \text{vbl}(A)$ do

$v_X \leftarrow$ a random evaluation of X

 EndFor

EndWhile

Return $(v_X)_{X \in \mathcal{X}}$

- $\text{vbl}(A) \subset \mathcal{Y}$: the set of variables determining A

Directions of LLL research

- Local conditions
 - Cluster LLL
 - Random walk
- Algorithms (Inspired by [Moser&Tardos](#))
 - Efficient beyond Shearer's bound?
 - Efficient for abstract events?

Comparing probabilistic methods

- All dependent vs almost independent
 - Counting (union bound): mutually exclusive
 - First moment: linearity doesn't care dependence
 - Second moment: pairwise dependence
 - LLL: global dependence

References

- Spencer. Ramsey's theorem-A new lower bound. 1975
- Spencer. Asymptotic lower bounds for Ramsey functions. 1977
- James B. Shearer. On a Problem of Spencer. 1985
- Robin Moser and Gabor Tardos. A constructive proof of the general Lovasz Local Lemma. 2009
- Polipaka and Szegidy. Moser and Tardos Meet Lovász. 2011
- <http://www.openproblemgarden.org/>

Thank you